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A Study of Genetic Variability in Larval and Adult Populations of Dungeness Crab (Cancer Magister)

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A STUDY OF GENETIC VARIABILITY IN LARVAL AND ADULT POPULATIONS
OF DUNGENESS CRAB (*CANCER MAGISTER*)

A Thesis

Presented to

The Faculty of the Department of Biological Sciences

San José State University

In Partial Fulfillment

of the Requirements for the Degree

Master of Science

by

Cairbre Fanslow

August 2013

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The Designated Thesis Committee Approves the Thesis Titled

A STUDY OF GENETIC VARIABILITY IN LARVAL AND ADULT POPULATIONS
OF DUNGENESS CRAB (*CANCER MAGISTER*)

by

Cairbre Fanslow

APPROVED FOR THE DEPARTMENT OF BIOLOGICAL SCIENCES
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August 2013

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ABSTRACT

A STUDY OF GENETIC VARIABILITY IN LARVAL AND ADULT POPULATIONS OF DUNGENESS CRAB (*CANCER MAGISTER*)

by

Cairbre Fanslow

Spatial and temporal genetic variation was assessed using mitochondrial DNA from adult and megalopae of Dungeness crabs, *Cancer magister*, from Fraser Delta in British Columbia, Washington, Oregon, and California. Megalopae were continuously sampled using light traps in one location, Coos Bay, Oregon, and compared with adults sampled across the species' range. While small-scale spatial and temporal variation were evident, there was no significant fit to the Isolation-by-Distance model when considering adult and larvae populations sampled over a coastal length of ~1700 km (Mantel $r = 0.059$, $p = 0.289$). Smaller bays showed higher genetic diversity than larger bays, contrary to typical expectations. MtDNA diversity measures (assessed by canonical correlation analysis) were correlated with bay size and depth, suggesting that hydrological patterns or fishing harvests in local areas (or a combination of both) influenced the observed genetic signature. The Dungeness crab ancestral population likely went through a historic population expansion associated with recent glacial relaxation, as evidenced by mismatch distributions and predominance of singleton haplotypes. In larval samples, we commonly observed numerous individuals of a haplotype, which suggests that families of larvae remain aggregated during dispersal. Sampled offspring revealed a low fraction of the adult gene pool (in accordance with the Hedgecock effect), yet high variability in recruitment over time. A realized long distance dispersal of larvae and the additive effects of recruitment from different sources over time may have prevented significant genetic differentiation within the *C. magister* range.

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INTRODUCTION

Determining accurate dispersal scales and other mechanisms that affect population structure is essential for establishing policy management in marine environments (Palumbi 2003, Botsford et al. 2003, Trakhtenbrot et al. 2005). The question of whether marine environments are open systems, interconnected systems, or closed systems in terms of gene flow is still debated (Cowen et al. 2000, Warner & Cowen 2002, Sotka et al. 2004). Relevance of whether or not marine environments are open helps shape marine reserve design (Palumbi 2003) and helps determine the impact on fisheries and stock management. Understanding the dispersal of the juvenile larvae is critical to the question of marine conservation. Tracking larvae directly is not usually possible and populations are more often analyzed by observing the recruitment stage (Shanks et al. 2003), by modeling of passive particles (Kinlan & Gaines 2003, Sotka et al. 2004, Kinlan et al. 2005), or by inference using genetic markers (Palumbi 2003, Hellberg et al. 2001). Combining genetic diversity measures with ecological and biogeographic measures has been more recently established in the field of landscape genetics (Manel et al. 2003), and now “seascape” genetic studies (Galindo et al. 2006, Selkoe et al. 2008).

Predicting the realized gene flow in species with long pelagic larval durations is particularly difficult. The duration larvae spend in the water column and the potential dispersal distances are likely important factors in the genetic structure of marine organisms. The major method of understanding larval dispersal, and thereby population connectivity, is through genetic markers (e.g., Wright 1943, Weir & Cockerham 1984,

Avise et al. 1987, Warner & Cowen 2002). The use of genetic markers to establish phylogeographically patterns of gene flow within a species' geographic range has a long history (Avise et al. 1987, Hellberg 2009). Over time, genetic changes have accumulated, indicating complex patterns of dispersal more than passive drifting could explain (Palumbi 1994, Galindo et al. 2010). As evidence of complex dispersal has accumulated, genetic markers in species with long planktonic duration have often shown genetic breaks occur over short distances (Kelly & Palumbi 2010).

Larval behavior, perhaps as importantly as pelagic duration, is an important aspect of population dynamics. Active swimming behaviors of larvae, such as vertical migration, are the primary means by which many species avoid some tidal actions or rapid surface currents (McLaren 1963, and for *C. magister*, Shanks 1986, Pillar et al. 1989). Oceanographic features, such as tidal forces and upwellings, have been correlated with larval distribution (Shanks et al. 2003). Tidal and upwelling correlations can then be used to provide spatial models to estimate larval distribution. In predicting population dynamics, further biological and oceanographic factors need integration. Factors that need to be considered include the fecundities of organisms and ecosystem productivity, which in at any given time affect spatial densities of recruiting populations and seasonal or yearly variability in current pattern (Shanks et al. 2003; Cowen & Sponaugle 2009).

As an example of important marine geographic barriers, Wing et al. (1995) revealed how currents around Pt. Reyes and Pt. Arena, California cause offshore accumulation of crab larvae during seasonal upwelling; after accumulation - during upwelling relaxation events - larvae then tend to be transported back to the vicinity of

origin. Point Conception in Southern California is a one of the most highly studied biogeographic barriers (Burton 1998, Dawson 2001). Here, numerous case studies reveal diverse evolutionary dynamics. There are cases of species in which dispersal breaks occur at Point Conception and cases of ecologically similar species showing genetic connection across this zone (Palumbi 1994, Hellberg 1996, Burton 1998, Dawson 2001). In a study of three limpet species over shorelines of multiple Hawaiian islands, Bird et al. (2007) found stronger evidence of zonal genetically segregation across intertidal-sub tidal depth zones than between islands. Case studies, such as those described by Bowen et al. (2006), indicate that pelagic duration correlated poorly with genetic subdivision and that the importance of ecological variables may determine genetic structure.

A number of studies have countered the simple “rule” that long distance dispersal should predictably lead to genetic cohesion. Johnson and Black (1982) found a shifting, “chaotic” genetic patchiness among planktonic dispersing limpets from Western Australia. This reported genetic patchiness found in allozyme markers was one of the first examples of the unevenness of population structure in an apparently open marine system. Doherty et al. (1995) also found that genetic variability among seven fish species sampled throughout the Great Barrier Reef that did not match an Isolation-by-Distance model. Sotka et al. (2004) found that acorn barnacles in the rocky intertidal communities of the Pacific Ocean show strong population division in the region of central California but not in Southern California, or Northern California, or Oregon. As a whole, population differentiation can be influenced by a mixture of larval behavior, ecological

impact, and especially reproductive success, all of which contribute to the unique genetic signature in long-distance dispersal species.

It is generally not expected that all individuals in a population contribute equally to reproduction simultaneously. Temporal variation in larval success, where a small fraction of the reproducing adult population is represented each year due to oceanic variability and high larval mortality, may explain genetic patterns present within a species at any given moment. Hedgecock (1994), considering demographics in terms of effective population size, N_e , and census population size, N , illuminated the point that a relatively small proportion of adults may account for the bulk of reproduction at one time, particularly in broadcast spawners. The process also envisions organisms as having relatively few “winners,” that is, successful breeders, and many “losers” in reproduction (Hedgecock & Pudovkin 2011). The idea has gained interest, and a “sweepstake” reproduction pattern is often referenced to describe genetic variance in highly fecund marine species (Li & Hedgecock 1998, Turner et al. 1999, 2002, Flowers et al. 2002).

In systems in which genetic variance in the gamete or larval pools vary strongly over time, seasonal variations are likely to be important in determining the genetic variance accruing in and among populations. The “sweepstake theory” predicts that species, though they may be broadly dispersed, are not necessarily highly similar in genetic structure.

In agreement with the sweepstake idea, Turner et al. (1999) and Flowers et al. (2002) both found that recruitment-level variation impacted population structure, even overcoming small selective advantages and disadvantages (fitness variation) of alleles.

Flowers et al. (2002) researched variance in reproductive success contributing to each larval cohort in the sea urchin, *Strongylocentrotus purpuratus*, evaluating 348 mitochondrial sequences. They observed high temporal genetic variance, as larval populations were genetically and statistically distinguishable when compared across years, and estimates of N_e , the effective number of breeding females, were generally low compared to the census population size. Similarly, in a study of Red Drum, *Sciaenops ocellatus*, an estuarine fish, Turner et al. (1999) found evidence of genetic drift changing the genetic signature in different years (1986–89), from which they determined that the effective (female) population size was small in each year.

As Turner et al. (1999) showed, temporal variance in the N_e can be quite high between years. These studies emphasize that assumed genetic variance is temporarily stochastic. Models that assume an equilibrium N_e measure (such as the gene flow-drift equilibria model) would be unrealistic in species with yearly variance in reproductive conditions. The question remains, can annually changing oceanographic and reproductive conditions explain more of genetic variance in broad spawning species than a genetic model based on distance?

Research system and goals

The goal of this thesis was to compare populations of the Dungeness crab, *Cancer magister*, in space and over time using cytochrome c oxidase subunit 1 (COI) nucleotide sequences to determine how population diversity varies along the coastline and to assess whether there are genetic discontinuities that are relevant to stock management.

Cancer magister, which is sometimes referred to taxonomically to the genus *Metacarcinus* (Ng et al. 2008), occurs along the western Pacific coastline of North America from the Aleutian Islands, Alaska, to Baja California, Mexico. The commercial range extends from Alaska through the Puget Sound, south along Washington and Oregon, ending in Southern California (Jensen & Armstrong 1987, Pauley et al. 1989). *C. magister* is common at depths of <100 m along the shelf, occurs within shallow marine estuaries on the North American Pacific coastline, and occurs at lower density in Southern California. Wild (1980) suggested that low *C. magister* abundance in Southern and Baja California is a consequence of poor adaptation to higher water temperatures.

Detection of genetic patterns through genetic markers has generally increased understanding of population differentiation. The COI gene in particular has been used frequently in intra-specific genetic variation studies (Avise et al. 1987, Barber et al. 2002, Waples et al. 2008, Bucklin et al. 2011). Population connectivity can be tracked due to the similarity of haplotypes defined by the (usually silent) nucleotide substitutions seen in this mitochondrial DNA segment. Examples of genetic differentiation using the genetic marker COI include the Eurasian perch, *Perca fluviatilis*, (Bergek and Olsson 2009); copepods along the California coastline, *Tigriopus californicus*, (Burton 1998); the barnacle *Balanus glandula* within rocky intertidal zones (Sotka et al. 2004); and in a wide range of marine species along the Western coastline of North American (see Kelly & Palumbi 2010).

Utilizing collections of adult and megalopae stage crabs and COI analysis, this thesis examined spatial gene flow over large and small scales and secondly, annual variance in larval populations was examined, comparing 6 (non-consecutive) years of larvae collected at a range of locations, with intensive sampling at Coos Bay, Oregon.

It is still generally difficult to predict whether gradual patterns of genetic divergence (isolation-by-distance), or sharp, barrier-associated disjunctions are present in a given organism's range. In the context of the Pacific Ocean, it is relevant to ask whether long duration pelagic dispersers avoid barriers or whether there are more localized patterns influencing a species' genetic disjunction.

The Dungeness crab was selected as a model invertebrate to study larval dispersal because it has high fecundity, long reproductive life, and long-range larval dispersal. The crabs live for an average of 8 to 10 years and reach effective breeding age by 2 years (Pauley et al. 1989). Adult males (which embrace females during reproduction and hold fast for up to a week) covet premolt females. Females typically produce 2 million eggs, but production can reach an estimated 5 million eggs (Pauley et al. 1989, Fernandez et al. 1993). Egg release generally occurs generally between December and April, depending on latitudinal temperature variation (Wild 1980, Moloney et al. 1994). The long distance dispersal capability of Dungeness crab larvae (Shanks 2006, Shanks 2009) may reflect a general adaptation favored by oceanic variability. After eggs hatch and become planktonic, tidal forces carry the larvae out of bays and estuaries onto the continental shelf. *C. magister* larvae undergo up to 5 months of development in the open ocean

through five zoeal stages, with the megalopae being the last stage before molting into recognizable juveniles (Roegner et al. 2003). *C. magister* is a predator throughout its life cycle, feeding in the open ocean carnivorously on plankton and then as adults on bivalves, mollusks, fish, and even on smaller *C. magister* (Pauley et al. 1989).

After the transition from larvae to megalopae, the megalopae recruit back to estuaries and coastal sites along the Pacific Coast range. “Recruiting” cohorts are generally defined as those surviving after the zoeal development when megalopae are returning to various sites along the coast. Recruits are dispersed to the continental shelf by upwelling currents and atmospheric pressures acting on ocean currents and then back to shore by using the same hydrological processes (Shanks & Roegner 2007). *C. magister* megalopae recruitment may depend on specific temperatures before entering estuaries in high abundance from April through September (Pauley et al. 1989). Within a few months, megalopae arrive along the coast and enter estuaries in distinct pulses interspaced with periods of low or zero abundance (Fernandez et al. 1993, Roegner et al. 2003, Shanks & Roegner 2007). After recruitment to suitable habitats, megalopae develop into juveniles. Juveniles settle in <20-meter deep benthic habitats along the open coastline or within estuaries; however, the mechanisms for crab recruitment into estuaries are not well understood (Roegner et al. 2003).

Toonen and Grosberg (2003 unpublished: report to California SeaGrant) set out to determine whether *Cancer magister* showed geographic structure or whether there were differences in the haplotypes of recruiting sets of larvae. While many microsatellites were isolated, the authors did not find large enough populations of *C. magister* to study

effectively. In the same work, the authors reported findings for another species, the porcelain shore crab, *Petrolisthes cinctipes*, which indicated a lack of geographic spatial pattern. The authors also found that while there are differences in the genetic patterns of separately collected samples of larvae, the differences were dampened over time, as multiple sets of recruits arrive.

In a study of COI sequences from eight populations of a common shoreline crab, *Hemigrapsus oregonensis*, Petersen (2007) found differences indicating barriers to dispersal among the northern and southern regions, corresponding to Cape Blanco (approximately the California and Oregon state boundary). COI haplotypes overall had wide ranges of distribution, indicating that since post-glacial range formation, there has been long distance dispersal, yet some restriction across Cape Blanco was observed. Petersen referred to the Cape Blanco area as a “semi-permeable” barrier. The contribution of recent dispersal events (including patterns driven by El Niño patterns, for example) versus the effect of historic isolation patterns has not yet been quantified in many cases. Statistics that are sensitive to haplotype frequencies in combination with phylogeny, such as Tajima’s *D* and the mismatch distribution (explained below), showed distinguishable patterns in the northern and southern *H. oregonensis* populations. Northern populations had more endemic haplotypes, whereas the signature of the southern area was consistent with an entire population expansion into the area following the most recent ice age. *H. oregonensis* reportedly has a 3-4 week planktonic larval duration. Petersen argued that long-duration-larval (six week or more) species such as *C.*

magister may show no population structure and no evidence of ancestral refuges shaping diversity (Petersen 2007).

C. magister has been the subject of a number of genetic studies, including unpublished theses. While generally examining small regions of the range, studies have demonstrated homogeneity in regional scale analysis of population structure (Dedhia 2005 unpublished thesis, SJSU, Lardy 2006 unpublished thesis, SJSU, Beacham et al. 2008, Barney 2012 unpublished thesis, SJSU) found that genetic connection is generally high among populations sampled from California to Washington.

In a study of 10 Puget Sound sample locations, Beacham et al. (2008) sampled between 23 and 119 adult crabs per site and analyzed microsatellites. They found that the *C. magister* sampled around Victoria Island, British Columbia, were genetically homogenous, and only one sampled site, Alison Sound, had local differentiation (overall $F_{ST} = 0.031$; no p-value given). Beacham et al. (2008) speculated that narrowness and shallow depth of the channel leading into Alison Sound constrained the water volume exchange and accounted for a high level of larval retention in the single subdivided population. A test for isolation-by-distance revealed no discrete population division from Northern to Southern British Columbia despite diverging seasonal currents around Vancouver Island. Dedhia (2005) and Lardy (2006) studying adult *Cancer magister* populations both found high haplotype diversity (finding 350 unique haplotypes among 846 individual COI sequences) and low but significant genetic differentiation among populations. Barney (2012), using 445 sampled adults from British Columbia to Southern California, found that there was no significant population structure, or isolation-

by-distance pattern, but found a large number of site-specific mtDNA haplotypes; (of the 302 base pairs sequenced, gene diversity was relatively high, ranging from $h = 0.6$ to 0.8). This study, being the largest in terms of sample size and sampled range, found little restriction to gene flow.

The major goal of the present study was to examine population dynamics of *C. magister* by considering genetic analysis of larval samples and genetic patterns in adult population samples. Using COI sequences, we assessed the following questions below.

- (i) We examined collections of adults and tested for isolation-by-distance across the range.
- (ii) Larval samples collected during the spawning period of April through September in 2001, 2003, 2008, and 2009 in one location, Coos Bay, Oregon, were used to determine whether pools of recruits are genetically mixed populations or not.
- (iii) Analyses of genetic diversity measures were conducted to determine whether ocean currents or other localized factors (e.g., bay size, commercial catch) impact population genetic signatures.

MATERIALS AND METHODS

Sample collection

Adults. We obtained 296 adult *C. magister* samples from British Columbia, Washington, Oregon, and California (Figure 1), mostly from commercial catches, with a smaller proportion donated by local commercial fishermen, seafood markets/distributors, and scientists (Dedhia 2005, Lardy 2006, Barney 2012). Crabs were caught using standard circular crab traps (Dahlstrom & Wild 1983, Lardy 2006). Generally, a single leg was removed from each individual. In the case of live crabs, non-lethal sampling was employed. A fifth-walking leg was held firmly until the crab autotomized it. This approach minimizes the risk of excessive bleeding (Lardy 2006). In some cases, loose legs that had been lost by crabs while in traps or in holding tanks were included in the samples. To avoid accidental sample duplication, legs from the same source were sorted, and only those legs of the same size, from the same side of the body were used. Nevertheless, commercial fishermen stated that they rarely observed crabs missing more than one leg (Dr. L. Parr pers. comm.). In one case, over 200 legs and claws were donated from a single commercial fishing company in San Francisco, and a random sample of 53 was taken by using a random number generator (Lardy 2006).

One challenge to the methodology of adult sampling might be a lack of precise source location determination. The adult *C. magister* samples were non-randomly, opportunistically sampled through out the inlets illustrated in Figure 1. The sample locations included 3 potential areas: 1) the dockside of the nearest marina, 2) throughout

the bay, and 3) outside the bay on the open continental shelf, where local fishermen reported they had collected samples.

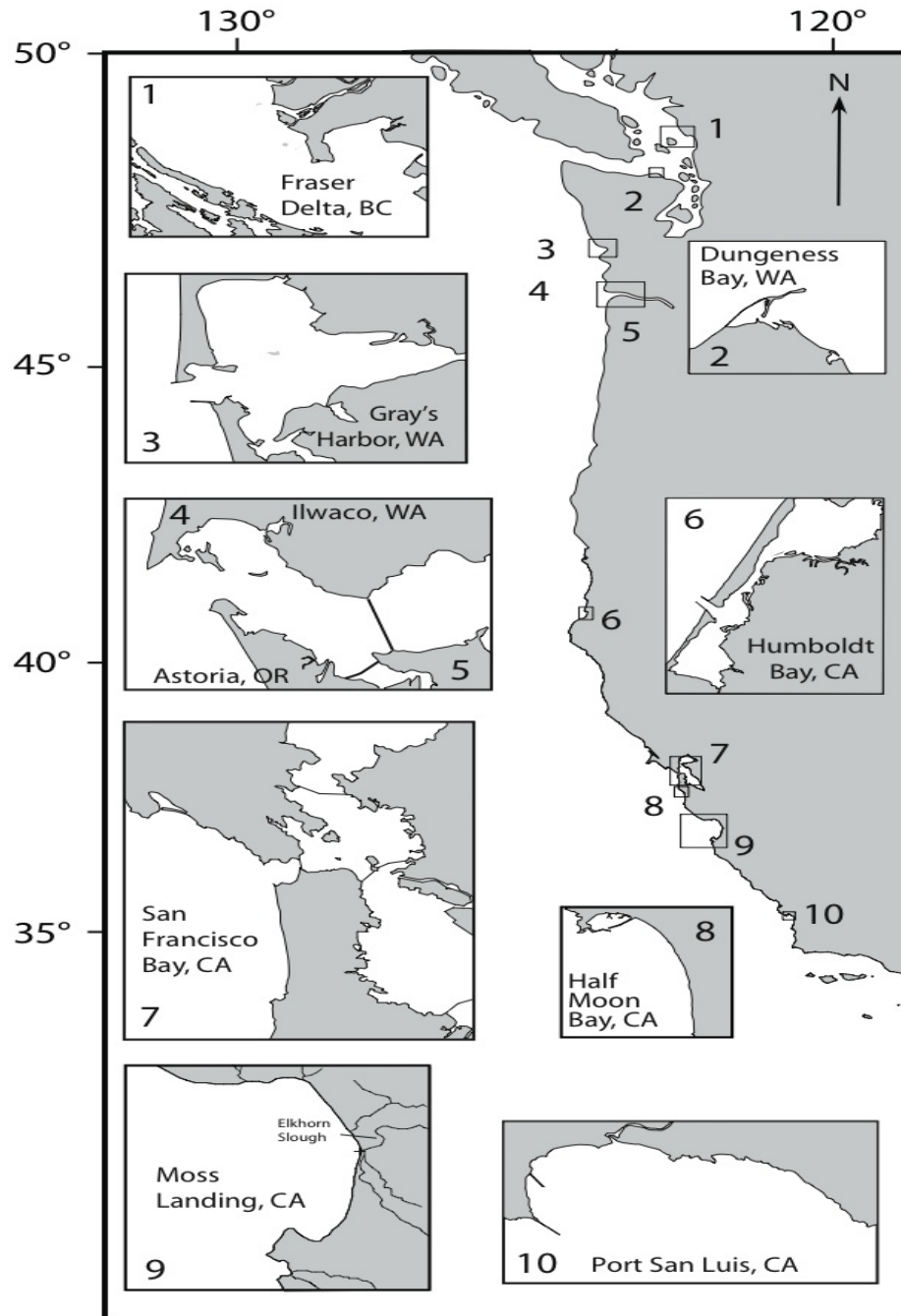


Figure 1. North American coastline with inserts showing the shoreline of sample sites. Graphics: Lynn McMasters.

Latitude and longitude coordinates were found using Google Earth 6.1, and distances were measured following a coastline trace at 100 feet altitude. In addition to COI sequence data reported here, sequences generated by previous graduate students (Dedhia 2005, Lardy 2006, Barney 2012) were compared to ascertain a more complete picture of Dungeness crab population structure.

Table 1. *Cancer magister* adult and larval population samples. Approximate geographic position < 2km; collector: Bryan Barney (B) or Alan Shanks (A).

Site	Year	n	GenBank accessions	Collector
Fraser Delta, British Columbia	2009	25	JX097359 – JX097384	B
Dungeness Bay, Washington	2009	24	JX097385 – JX097409	B
Gray's Harbor, Washington	2009	25	JX097409 – JX097434	B
Ilwaco, Washington	2009	25	JX097435 – JX097459	B
Astoria, Oregon	2004	24	JX097460 – JX097484	A
Eureka, California	2005	25	JX097485 – JX097510	A
San Francisco, California	2005	53	JX097511 – JX097564	A
Half Moon Bay, California	2005	36	JX097565 – JX097601	A
Moss Landing, California	2005	30	JX097602 – JX097632	A
Port San Luis, California	2005	29	JX097633 – JX097662	A
Coos Bay, Oregon	2001	135	JX097663 – JX097798	A
Coos Bay, Oregon	2003	60	JX097799 – JX097859	A
Coos Bay, Oregon	2008	169	JX097860 – JX098026	A
Coos Bay, Oregon	2009	33	JX098027 – JX098060	A

Megalopae. This study generated COI sequences from 406 megalopae-stage *C. magister*. Larval crabs were continuously sampled using light traps at the mouth of Coos Bay, exploiting the positive phototropism of megalopae towards artificial illumination (Miller & Shanks 2004, Roegner et al. 2007). Larvae were collected every few days and donated by Dr. Alan Shanks of the Oregon Institute of Marine Biology, Charleston, OR, from the Charleston Marina Coast Guard station pier in Coos Bay, OR, estuary at Long: 43°20'12"N by Lat: 124°19'14"W following the techniques outlined in Roegner et al. (2007). In total, 406 megalopae larvae of *C. magister* were sequenced for COI spanning 4 non-concurrent years. Years 2001 and 2003, were examined by previous graduates Dedhia (2005) and Lardy (2006) while sequences for megalopae collected in 2008 and 2009 were newly generated.

DNA extraction and PCR

Samples were held frozen (–20°C) for up to two years prior to DNA extraction and Polymerase Chain Reaction (PCR) amplification. DNA was extracted from leg muscle of adults, or from the whole organism for megalopae, using the Promega Wizard® Genomic Purification Kit Catalogue # A1125, following manufacturer's instructions. Extracted DNA was rehydrated in 50 µl of TE buffer and stored at –20°C.

PCR amplified a 359 base-pair amplicon, corresponding to positions 1718-1777 within the COI gene using primers designed specifically for *C. magister* (Harrison and

Crespi 1999). The forward primer was S1718b (5'-GGAGGATTTGGAAATTGATT-3'), and the reverse was 'AS2097 (5'-GTACAGGAAGGGATAGTAGT-3'). For statistical analysis, we used a common 338- base pair region for adults and a 222-base pair region for megalopae.

PCRs were conducted in 25- μ l volumes, containing 2 μ l of DNA dilution, 1X PCR Buffer (Finnzymes), 0.2 mM dNTPs (Qiagen), 0.2 μ M forward and reverse primers (Operon), 0.4 mg/ml bovine serum albumin (BSA) (Promega), 3.0 mM MgCl₂ (Promega), and 1 unit Taq polymerase (Finnzymes). The BSA was added to counteract the PCR-inhibiting affects of seawater (Kreader 1996). PCRs were run on a Personal Thermal Mastercycler (Eppendorf International, Hamburg, Germany): 5 min at 94°C, followed by a “touchdown” steps of 10 cycles of 30s at 94°C and 45s at 55°C, reducing by 0.5°C every cycle, 1 min at 72°C, then another 30 cycles of 30s at 94°C, 45s at 50°C, and 1 min at 72°C, and a final extension of 72°C for 5 min. Products were checked electrophoretically on a 2% ethidium-labeled agarose gels to check for specificity of amplification.

Characterization and submission of sequence data

PCR products were cleaned following instructions in Qiagen QIA Quick PCR purification kit #51504. Purified PCR products were sent to Geneway, LLC, at Hayward, CA for sequencing. Sequencing was conducted on an ABI Prism Model 377 (version 3.2) automated sequencer using Big Dye chemistry (Applied Biosystems). Sequences

were aligned and verified for accuracy of trace files and the removal of ambiguous bases, using Mega4 (Tamura et al. 2007).

Molecular data indices were calculated for population samples using the software Arlequin 3.5 (Excoffier & Lischer 2010). Diversity measurements included gene diversity (h), which is the probability that two random sequences taken from the same population will differ by at least one nucleotide, and nucleotide diversity (π), the average genetic distance between pairs of sequences (Nei 1986). Kimura-2 parameter distance was used in π , correcting for multiple substitutions per site while taking into account different rates or occurrence of transitions and transversion mutations (Kimura 1980). Once sequenced, all COI data sets were submitted to Genbank and received accession numbers JX097359 to JX097662 for adults and accessions JX097663 to JX098060 for megalopae (Table 1).

Neutrality tests

Fu's F_S (Fu 1997) and Tajima's D (Tajima 1989) are techniques to detect departures from neutrality or deviations due possibly to historic change in population size (groups of observations that themselves may not be distinguishable). Fu's F_S statistic indicates the presence of an excessive number of rarer and singleton alleles in populations. Negative scores of Fu's F_S test may signal population expansion (Fu 1997). Tajima's D scores are used to test the assumption of selective neutrality of mutations among COI sequences (Tajima 1989). In the absence of strong selection and with a historically, stable population size and a constant rate of mutation, D and F_S have

predictable equilibrium values. A test of significance for the D and F_S values were calculated from the infinite allele equilibrium relationship following Fu and Li's (1993) methodology outlined in Excoffier and Lischer article (2010). Populations that have gone through a bottleneck or founder effect in their recent past have characteristically negative D and F_S values (Mousset et al. 2004). Given results in a similar species in this region (Petersen 2007), we expect that Tajima's D and Fu's F_S scores for *C. magister* will be negative, assuming the population is expanding possibly from a recent bottleneck after the most recent glacial relaxation.

The mismatch distribution – a metric sensitive to historic population expansion

The mismatch distributions (a histogram of the numbers of mutational differences between pairs of haplotypes) were also used to test whether mutational distances among alleles could fit the “sudden expansion model” (Rogers & Harpending 1992). The mismatch distribution was also used to calculate r , the “raggedness” index (Harpending 1994). A low “raggedness” index value (< 1) would indicate a rejection of the fit between the observed distribution and the statistical model. Here, a p-value was determined by the sum of square deviations (SSD) test between the observed and the expected mismatch as a test statistic model (see Excoffier & Lischer 2010).

The mismatch is usually multi-modal in samples drawn from populations at demographic equilibrium, as it reflects the highly stochastic shape of gene trees, while it is usually uni-modal in populations having passed through a recent demographic expansion (Slatkin 1991, Rogers & Harpending 1992). Typically, multi-modal mismatch

distributions have large values of r ; uni-modal wave distributions typically have low values (< 1) of r .

To compare estimates of when adult and larval populations underwent recent population expansions, we calculated mismatch distributions to compared rates of population expansion (Harpending 1994; Rogers 1995). For expanding populations, we converted the parameter tau (τ ; calculated from the mismatch distribution) estimating time since population expansion (t) via the equation $\tau = 2ut$, where $u = \mu k$, where μ is the mutation rate per site per year, and k is the sequence length (Rogers 1995; also see Schneider & Excoffier 1999). We estimated t using a mutation rate (μ) of 2.3% per million years times the COI sequence for adults and larvae, estimated for a sample of Decapoda (Knowlton & Keller 1993, Knowlton & Weigt 1998). The lower and upper 95%-confidence intervals of τ were calculated using a parametric bootstrap approach (Schneider & Excoffier 1999).

Measurement of population differentiation

Populations were assessed for genetic difference using pairwise F_{ST} and hierarchal ordering using an analysis of molecular variance (AMOVA) (Excoffier et al. 1992) in Arlequin 3.5. AMOVAs were used to determine if there were genetic differences among adult populations and then again for megalopae across sampled years. Genetic differences among individuals were estimated by pairwise F_{ST} statistics over all loci (Weir & Cockerham 1984, Slatkin 1985) from the COI aligned sequences. The fixation indices were estimated using Kimura's 2-parameter distances among sequences

(Kimura 1980). Significance was determined in Arlequin 3.5 using 10,100 replicates, where p is estimated as the proportion of randomized permutations, of F_{ST} larger or equal to the observed F_{ST} .

To determine which source of variance (geographic or temporal) impacts population structure, this thesis examined spatial gene flow over large and small scales in space then examined temporal variance in samples over 6 non-consecutive years (Table 2). A series of AMOVA comparisons were conducted with the goal of contrasting variances among the sampling units to determine which was greater - temporal or spatial variance.

Table 2. Regional groupings used in AMOVA. n enumerates crabs of adult (A) or megalopae (M) age class. Region (North, Central, or South) is the geographic region used in hierarchical comparisons.

Site	Year	Region	n	Latitude	Longitude
Fraser Delta, British Columbia	2009	North	25(A)	49° 4'59"N	123°14'39"W
Dungeness Bay, Washington	2009	North	24(A)	48° 9'41"N	123° 7'32"W
Gray's Harbor, Washington	2009	North	25(A)	46°55'10"N	124° 6'41"W
Ilwaco, Washington	2009	North	25(A)	46°17'60"N	124° 2'27"W
Astoria, Oregon	2004	Central	24(A)	47° 06'00"N	124° 20'00"W
Eureka, California	2005	South	25(A)	40° 47'00"N	124° 13'00"W
San Francisco, California	2005	South	53(A)	37° 27'00"N	122° 34'00"W
Half Moon Bay, California	2005	South	36(A)	37°29'00"N	122° 29'07"W
Moss Landing, California	2005	South	30(A)	36°48'00"N	121°47'00"W
Port San Luis, California	2005	South	29(A)	35° 80'00"N	120° 44'00"W
Coos Bay, Oregon	2001	Central	135(M)	43°20'12"N	124°19'14"W
Coos Bay, Oregon	2003	Central	60(M)	43°20'12"N	124°19'14"W
Coos Bay, Oregon	2008	Central	169(M)	43°20'12"N	124°19'14"W
Coos Bay, Oregon	2009	Central	33(M)	43°20'12"N	124°19'14"W

Tests for spatial differentiation: Exact and Mantel tests

An Exact test (Raymond & Rousset 1995) was performed to test for non-random distribution of haplotypes into population samples under the hypothesis of panmixia. Arlequin 3.5 haplotypic matrices tables were used to compare all pairs of data from the complete data set of 296 adults and 406 megalopae. Sample haplotypic data table was built using 10,000 random sample haplotype frequencies from the data set (Raymond & Rousset 1995). Observed and sampled haplotypic tables are compared for significance following the procedure. A significant score indicates that data is not randomly distributed, which may indicate differentiation across the range of *C. magister*.

A Mantel test (Mantel 1967) was conducted using Arlequin 3.5 to determine if there was a relationship between geographic distance and genetic differentiation. The first matrix consisted of geographic distance (km) between each pair of sample sites. The second matrix consisted of the pairwise F_{ST} differentiation between COI sequences between all possible pairs of sites. The test of data significance used 10,000 permutations. A significant correlation between geographic distance and differentiation among the COI sequences would indicate Isolation-by-Distance, whereas a non-significant correlation would indicate that propagules were mixed across the geographic range.

Marine environment variables

In addition to considering the effect of spatial distance among populations, the present study attempted to assess whether patterns in haplotype frequency were associated with fine-scale habitat geometric or environmental variables. The relationship

between simple genetic diversity of population samples (measured as h or π and other more complex measures including D and F_S and pairwise F_{ST}) was considered in correlation-based analyses with hydrological measurements including commercial catch numbers for the sampled years.

Length and width of embayment openings were measured using Google Earth 6.1 at an altitude of 100 km. The length and the width (LxW) of entry distances into embayments (a term used in the U.S. Bureau of the Census 1990 report) were measured. The span across the waterway from an outside source of water flow was calculated as length while the width was measured as a perpendicular line towards the coastline (see Figure 2). The LxW measurement procedure is shown in Figure 2A and 2B.

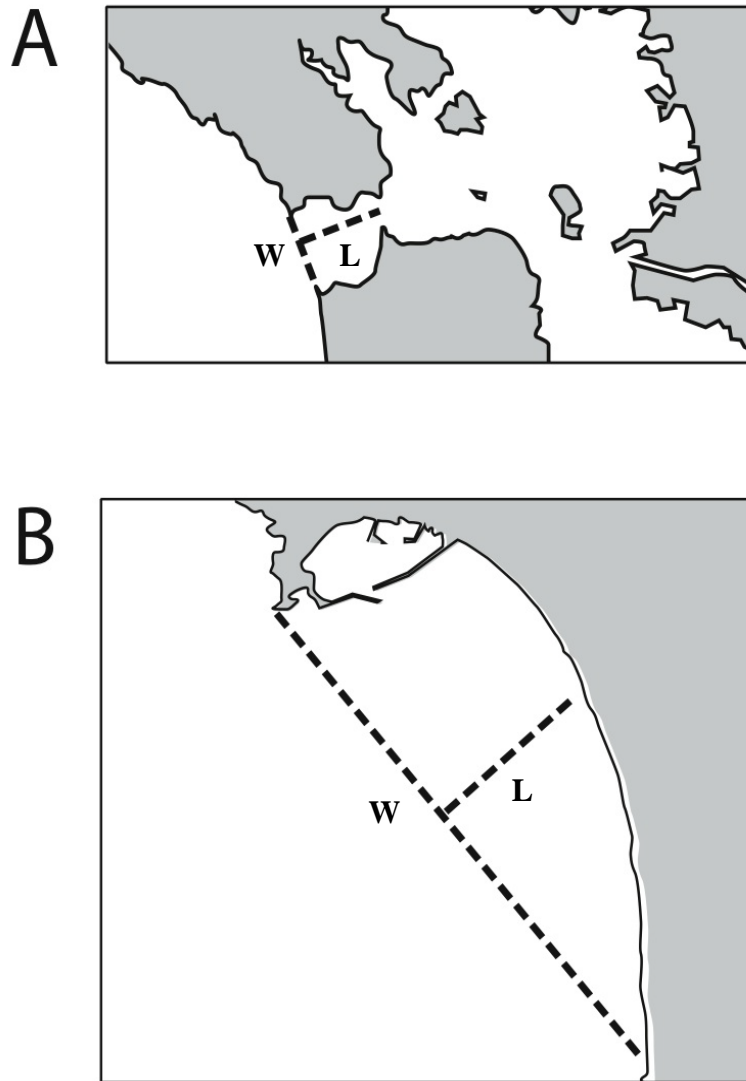


Figure 2. Schematic of embayments, showing a description of how bay opening areas by length x width measurements, were made (image is not to scale). A closure line was drawn from the headlands of embayments (at the western most points, in these cases). A line perpendicular to this length line was then extended to an inner edge of the inlet area as the width line. In the case of embayments that have narrow openings, such as San Francisco Bay (A) or Humboldt Bay (not shown), the perpendicular lines were extended to approximately the end of the inlet channel, as shown in (A). In the case of open-crescent shaped bays such as Half Moon Bay (B), the perpendicular line was extended to the coast. Accordingly, the bay area measurement approximates in km^2 the entryway that has unhindered contact with the open ocean.

The bay width line was extended to either the point at which the opening expands, in the case of enclosed structures, or to the point of the beach line, in the case of crescent shaped (or opening size of) bays. For example, measurement of opening size applied to the San Francisco Bay would extend from a closure line west of the Golden Gate Bridge eastwards to the shoreline to the narrowest point before the channel opens up into the SF bay, at which point, the water becomes classified as inland delta water (Figure 2A). As an example of an open-crescent shaped bay, the LxW measurement scheme used to define Half Moon Bay is shown in Figure 2B. The LxW measures approximate (in km²) the water domain that has unhindered contact with the open ocean and tabulated in Table 8. We confirmed that adult crabs providing data corresponding to bays were caught near (< 2 km) these embayments though a precise location was not possible to determine as explained above.

Mean depth (m) of embayments (Table 8) were collected from National Oceanic and Atmospheric Administration (NOAA) charts available online.

The average yearly commercial catch where available was reported (in tons [equivalent to 2,000 pounds] Table 8). Measurements were found online at state reported commercial catch numbers, which correspond to the years adult crabs were sampled. We expected that high catch areas may show different diversity from low catch areas, and plotted typical seasonal tonnage data with genetic diversity measures. All Dungeness crab fisheries are managed primarily by a minimum size limit and non-retention of females (Yonis 2010). For Fraser Delta, crab catch was reported by DFO (Department of

Fisheries and Oceans Canada) for 2009 and available online. Data for Washington State combined Dungeness crab, including Grays Harbor and Ilwaco, was reported for the 2009 season by Washington Department of Fish and Wildlife (WDFW) and available online (WDFW 2013). For California Dungeness crabs in the Northern region, which includes Half Moon Bay and San Francisco, catch data was reported by California Department of Fish and Wildlife (DFW) and included catch for Eureka Bay and Fort Bragg available online (DFW 2013). California's Department of Fish and Wildlife reported commercial fishery catch for Dungeness crab as two main sections along the coastline: Northern California and Southern California (DFW 2013). Southern California fishing areas include other site such as Avila-Morro Bay, Monterey bay, while the Northern area group San Francisco with Bodega Bay. The Morro Bay and Monterey fisheries are minor compared to the Northern San Francisco-Bodega Bay fisheries. The commercial catch regulations require throwing back of females (Department of Fish and Wildlife 2013, Washington Department of Fish and Wildlife 2013). There were no distinction among sampled sites of Half Moon Bay and San Francisco. Similarly, Southern California catch, which included Moss Landing and San Luis Obispo did not distinguish among coast sites and data is grouped with other Southern California coastal cities (DFW 2013).

Turbidity of embayments for each location was approximated using Nephelometric Turbidity Units (NTU) and data were collected from various published papers (Table 8). NTU is a measure of waterborne light-scattering particles when a light beam is focused through water and reflected back to a detector. The resulting light scatter can be used as a measure of turbidity. In this study, turbidity data as NTUs were

obtained from published records that were made using calibrated a nephelometer (Myre & Shaw 2006). Other studies have shown that NTU in many cases approximates the flow of water through embayments (White 1994). Turbidity and total solids may also increase sharply during and immediately following a rainfall, a temperature increase (which results in oxygen-concentration decrease), or response to toxin levels (EPA 2009). As one example, a study found that agricultural field runoff in its watershed contributed about 22,500 tons of sediment per year to Weeks Bay, in Alabama (EPA 2009).

Averaged current velocity (cm^3/s) data were utilized from each site to approximate the flow of water in and out of embayments (Table 8). Velocity data were obtained from acoustic Doppler systems of the publically accessible National Oceanic and Atmospheric Administration (NOAA) National Data Bouy Center (NDBC) database. Average velocity was determined as the current speed collected from a single surface buoy over 24 hour period, quality-controlled, and reported to the website.

We correlated matrices of pairwise F_{ST} measures of adult samples from Results section (Table 5A) with embayment opening size measurements (Table 8) as well as with D and F_S measures (Table 3A) to examine whether hydrology correlated to population diversity.

Examination of the association between genetic and environmental variables

Canonical correlation analysis (CCA) was used to determine if there was a relationship between a suite of environmental measures and a suite of genetic diversity measures. The environmental measures included opening size, mean depth, NTU, and

velocity in Table 8. The genetic measures included h and π as well as neutrality scores D and F_S (Table 3A and 3B). Canonical loadings, correlations between the variables and the canonical scores for the suite of variables being measures, were used to interpret the canonical scores. Positive loading indicates that small values of canonical score correspond to small values of the variable, whereas a negative loading indicates that small values of canonical scores correspond to the high values of the variable.

Canonical correlation analysis requires a linear relationship between variables: therefore, prior to testing, we plotted all possible relationships among genetic variables (h and π) and neutrality scores (F_S and D) with environmental variables (LxW, mean depth, NTU, and velocity) to verify these were linear. The Pearson test for independence revealed that gene diversity (h) and nucleotide diversity (π) were not independent and could be used interchangeably. Similarly, F_S and D were not independent and could be used interchangeably for correlation tests. The correlation test was a novel approach to analyzing genetic diversity spatial patterns using easily obtained environmental variables.

Pairwise F_{ST} and LxW scores were compared to determine if a relationship existed (Results section: Figure 6). To further explain any findings of correlation or genetic difference, as pairwise F_{ST} scores, and the highest correlated physical data (which was preliminarily opening LxW) were cross compared using Robust Linear Regression of observed data. The Shapiro-Wilk test indicated data was normally distributed and also provided a test for significance among data sets. This regression method was designed to not be overly affected by violations of assumptions in the underlying data-generating

process. In contrast, the least squares estimates for regression models might have proved too robust (might have given too much weight) to data outliers (Wilcox 1998).

The Wilcoxon-signed rank test (1945) was performed to compare pairwise F_{ST} extracted from the population matrix to the environmental variables for given populations in order to determine if data sets follow normal distribution to avoid a false positive result. Helping to avoid a false positive result, nonparametric (the Wilcoxon test being an example) tests do not assume that the data conform to a particular probability distribution. The Wilcoxon test gives the value Z , the sum of signed ranks divided by square root times the sum of squared ranks (Wilcoxon 1945). A p-value is accepted if a probability test calculates that observed data is less than a statistically derived estimate of data level of $p \leq 0.05$.

RESULTS

Molecular diversity

Comparison of COI sequences of 359 base pairs from 296 *C. magister* adults yielded 55 polymorphic sites and 86 haplotypes (Table 3A). There was a ratio of 34 transitions to 23 transversions and a total of 57 substitutions. The three most common haplotypes in the adult data set occurred at frequencies of 29.0% ($n = 86$), 21.9% ($n = 65$), and 8.4% ($n = 25$). No other haplotypes occurred at frequencies greater than 2% ($k = 6$), and the proportion of singleton (Slatkin 1985) sequences occurring once in the data set was 40.5% ($n = 120$).

Table 3A. *C. magister* adult diversity statistics. Significant ($p < 0.05$) test results are shown in bold; * $p \leq 0.01$; ** $p \leq 0.001$.

Population, (abbrev.), collection date	<i>n</i>	Number of private Haplotypes (<i>k</i>)	Gene Diversity (<i>h</i>)	Nucleotide Diversity (π)	Tajima's <i>D</i>	Fu's <i>F_s</i>
Fraser Delta (FD) 2009	25	8	0.630	0.003	-1.720	-3.902*
Dungeness Bay (DB) 2009	24	9	0.866	0.008	0.905	-1.541
Gray's Harbor (GH) 2009	25	9	0.823	0.007	-0.227	-2.295
Ilwaco (IW) 2009	25	9	0.817	0.004	-0.868	-4.435**
Astoria (AS) 2004	24	22	0.981	0.014	-1.007	-19.629**
Eureka (EK) 2005	25	14	0.947	0.008	-0.589	-7.550**
Half Moon Bay (HM) 2005	36	12	0.789	0.005	-1.322	-6.540**
San Francisco (SF) 2005	53	18	0.837	0.006	-1.875*	-10.507**
Moss Landing (ML) 2005	30	12	0.862	0.006	-1.183	-5.728**
Port San Luis (SL) 2005	29	14	0.857	0.005	-1.777	-10.139**
Total adults	296	90	0.851	0.007	-2.145**	-26.733**

Adults sharing haplotypes at one or more polymorphic sites constituted 68% of the population overall, while private haplotypes not shared with any other population constituted 32%. There was a ratio of 85 transitions to 48 transversions and a total of 133 substitutions. There were six synonymous changes in the first codon, and the remaining 42 synonymous changes (silent substitutions) were in the 3rd codon position.

Truncating the 2001, 2003, 2008, and 2009 megalopae dataset in order to include the 222 base pairs common with the previous megalopae dataset studies (Dedhia 2005, Lardy 2006) yielded 28 unique haplotypes (Table 3B). Truncation was necessary in order to ascertain the most precise reading of some sequences from previous megalopae datasets (2001 and 2003). The three most common haplotypes in the larval data set, which were the same three common haplotypes in the adult set, occurred at frequencies of 23.9% for haplotype 1 (n = 97), 24.9% for haplotype 2 (n = 101), and 8.62% (n = 35).

Table 3B. *C. magister* megalopae diversity statistics, pooling larvae collected in 2001, 2003, 2008, and 2009 at Coos Bay. Significant ($p < 0.05$) test results are shown in bold; * $p \leq 0.01$.

Collection year	<i>n</i>	Number of private Haplotypes (<i>k</i>)	Gene Diversity (<i>h</i>)	Nucleotide Diversity (π)	Tajima's <i>D</i>	Fu's <i>F_s</i>
Coos Bay 2001	135	4	0.583	0.006	1.385	2.580
Coos Bay 2003	75	8	0.801	0.016	0.887	2.572
Coos Bay 2008	165	14	0.680	0.004	-1.361	-8.109*
Coos Bay 2009	31	11	0.821	0.014	-1.697	-2.226
Total megalopae	406	28	0.829	0.010	-0.197	-1.296

Some larval samples had little or no diversity (Table 4). For example, only one haplotype was found in the light trap (which was deployed continuously since April 2001 and checked daily) for the July and September 2001 samples. When multiple weeks samples from 2008 and 2009 were considered (each week having sample-sizes of 8 to 81 megalopae per trap event) only two haplotypes were found in August 2008. The remaining sampled weeks had only one haplotype from April 2008 until August 2009 (see Table 4). The haplotypes tended to be common haplotypes (haplotype assigned number 1, 2, or 3), suggesting very few females contributed to the total brood population from Coos Bay. Exact contribution of females, in terms of numbers, would require use of a more variable class of genetic markers or more loci.

Table 4. Distribution of haplotypes found in megalopae populations. Megalopae were collected by light trap in Coos Bay, OR. Dr Alan Shanks oversaw collections.

Period	Approximate collection period (days)	Number of individuals analyzed	Number of haplotypes	Actual haplotypes present and frequencies (n)
July Wk4 2001	7+	12	1	5(12)
Aug Wk1 2001	7	14	1	3(14)
Aug Wk2 2001	7	28	1	1(28)
Sept Wk2 2001	7+	81	1	6(81)
Jun Wk2 2003	7+	9	1	1(9)
Jun Wk4 2003	14	24	2	1(1)
				7(23)
				14(1)
Jul Wk1 2003	7	21	2	8(20)
				9(1)
Jul Wk2 2003	7	6	2	10(5)
				11(1)
Jul Wk3 2003	21	15	2	2(14)
				15(2)
Aug Wk1 2003	14	17	1	12(12)
Aug Wk2 2003	7	10	3	9(1)
				13(8)
				14(1)
Apr Wk4 2008	7	8	1	1(8)
May Wk2 2008	14	8	1	1(8)
May Wk3 2008	7	12	1	2(12)
May Wk4 2008	7	11	1	2(11)
Jun Wk2 2008	14	11	1	2(11)
Jun Wk3 2008	7	12	1	2(12)
Jun Wk4 2008	7	12	1	1(12)
Jul Wk1 2008	7	11	1	2(11)
Jul Wk2 2008	7	10	1	1(10)
Jul Wk3 2008	7	12	1	1(12)
Jul Wk4 2008	7	12	1	1(12)
Aug Wk2 2008	14	10	1	2(10)
Aug Wk3 2008	7	12	2	1(11)
				2(1)
Aug Wk4 2008	7	12	1	1(12)
Sep Wk1 2008	7	12	1	3(12)
Apr Wk4 2009	7	10	1	3(10)
Jul Wk2 2009	7	10	1	1(10)
Aug Wk4 2009	7	11	1	4(11)

Population diversity

Nucleotide diversity (π) and gene diversity (h) of all sampled adults were similar to larval population total sampled diversity scores (Table 3A and 3B). Nucleotide and gene diversity for the entire adult dataset were 0.010 and 0.850 (Table 3A). For larvae from Coos Bay sampled over the years, (π) and (h) diversities had a net average of 0.010 and 0.829 (Table 3B).

Megalopae samples collected in 2001 and in 2008 had lower diversity on average than adults: for 2001 $\pi = 0.006$, $h = 0.583$; for 2008 $\pi = 0.005$, $h = 0.680$ (Table 3B). Megalopae sampled from 2003 and 2009 had higher diversity: for 2003 $\pi = 0.016$, $h = 0.801$; for 2009 $\pi = 0.014$, $h = 0.821$ (Table 3B), which were in the range of those generally seen in adult populations (Table 3A). Assessed over the time frame of a year cohorts of megalopae from Coos Bay, varied in diversity, h ranged from 0.63 to 0.98 and their nucleotide diversity, π ranged from 0.003 to 0.014 (Table 3B). When larvae were broken down by one-week time frames (Table 4) single haplotypes appeared over approximately 7 or more days of sampling, the repeated appearance of single haplotypes suggesting broods of larvae were collected together.

Figure 3 illustrates the diversity of adult populations in relation to their geographic distribution and embayment dimensions. As seen in the insert, narrow mouthed embayments had higher haplotypic diversity generally. Typically, three common haplotypes were found in each location (Figure 3).

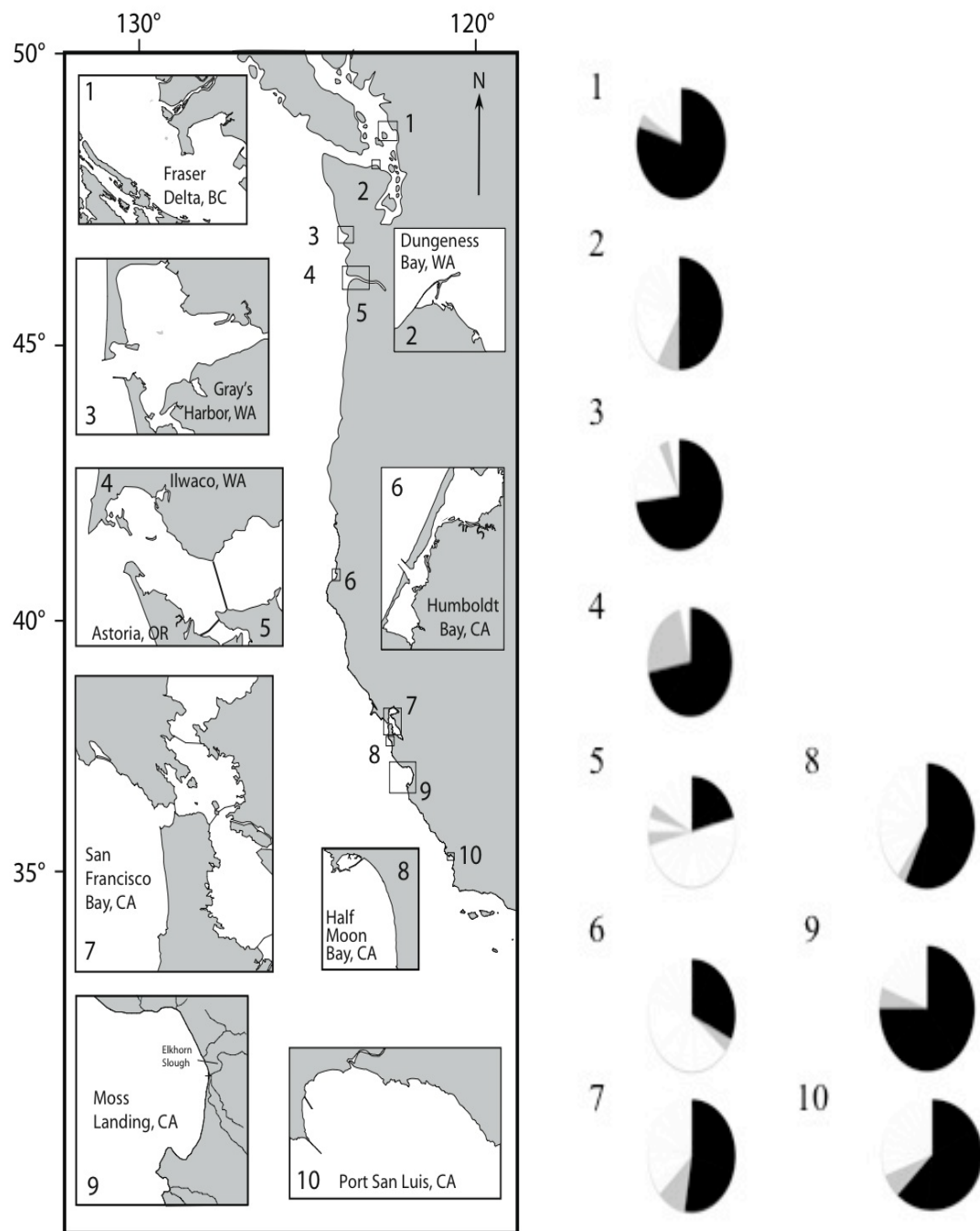


Figure 3. Diversity pie charts and collection inserts of embayments. Black chart segments represent frequencies of haplotypes found at all sites, grey indicates haplotypes shared at one or more sites, and white indicates private haplotypes found only at one site.

The black shading in pie charts (Figure 3) are shared haplotypes throughout the Dungeness crab's range, whereas grey shading represents haplotypes shared in one or more locations but not in all populations, and white portions are private haplotypes unique to that embayment. Throughout the Northern American Pacific range, each population site shares haplotypes 1, 2, and/or 3. As Figure 3 indicates, narrow mouthed bays like Dungeness Bay and Humboldt Bay have over 50% of private haplotypes, suggesting areas of high genetic diversity.

There are notable anomalies. At Astoria where over 50% are a private haplotype, but one km north on the same river is Ilwaco over 90% are shared geographically. Populations at San Francisco, Half Moon Bay, Moss Landing, and the port of San Luis Obispo have 50% or more shared haplotypes, which corresponds to the lack of regional differentiation found in the pairwise F_{ST} scores and AMOVA results. There was one notable exception to net diversity scores in the San Francisco population. In the San Francisco population ($n = 53$), haplotype 2 was absent, though haplotype 2 appeared in all other populations.

Neutrality test results

Tajima's D in adult population samples was generally negative in value, but only three of the ten populations showed statistically significant deviations from the neutral model (Table 3A). The remaining seven adult populations had non-significant p-values. Tajima's D was estimated in four larval samples from years of collections in Coos Bay

(Table 3B); two years had positive D values (2001 and 2003), and two negative D values (2008, 2009), and only one adult population was statistically significant (2009). A broad survey of 21 spp. of crustaceans suggests slightly negative D is the most common pattern (Kelly & Palumbi 2010).

As suggested earlier, F_S scores are particularly sensitive to detecting data sets that have a “star phylogeny” and the large numbers of alleles may indicate an historic population expansion (Fu 1997). In the sampled adult populations, F_S values were slightly negative, and all populations were statistically significant with two exceptions: Dungeness Bay ($F_S = -1.540$, $p = 0.212$) and Gray’s Harbor ($F_S = -2.295$, $p = 0.098$). Megalopae samples from Coos Bay (Table 3B) had two nominally negative F_S scores (2008 and 2009) but only one statistically significant score (negative F_S in 2008).

There was no obvious relationship in comparing neutrality scores and numbers of private haplotypes, which ranged from 8 to 22 private haplotypes in adult samples (Table 3A) and 4 to 14 in the megalopae sampled over several years at Coos Bay (Table 3B).

The mismatch distribution of both the adult and larval populations sets was a unimodal, smooth wave (Figure 4A, Figure 4B). The least squares test of the fit with the Sudden Expansion model (Schneider & Excoffier 1999) was non-significant, indicating distributions agree with a history of recent population expansion: for adults, the mean Sum of Squares deviations (SSD) was 0.008, $p = 0.469$, and for larvae SSD was 0.049, $p = 0.068$. Harpending’s raggedness index for the total adult populations set was $r = 0.062$, $p = 0.458$, and for the larvae populations was $r = 0.172$, $p = 0.051$.

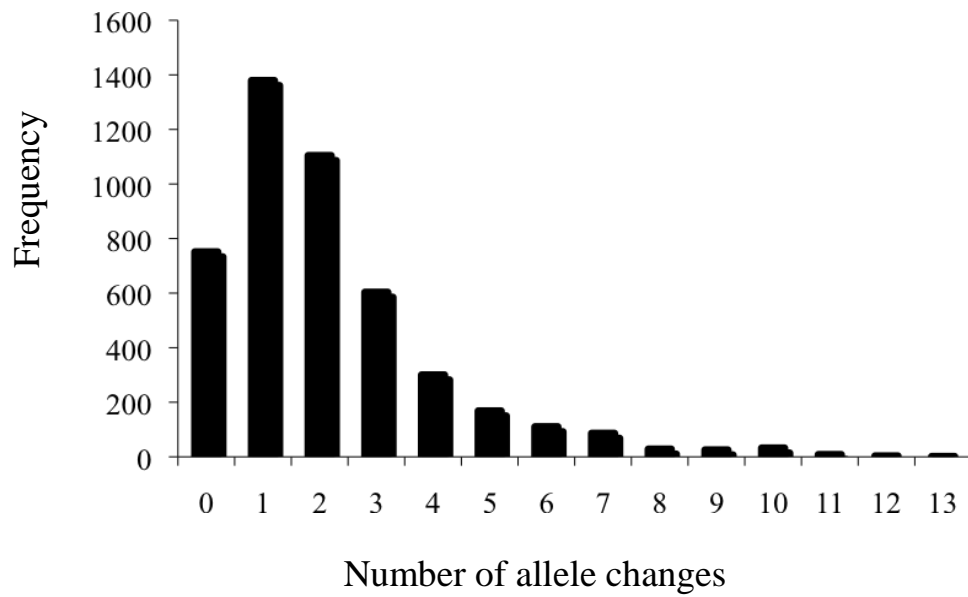


Figure 4A. The mismatch distribution of COI sequences obtained from adult populations. The mean tau value of allele changes was 2.253.

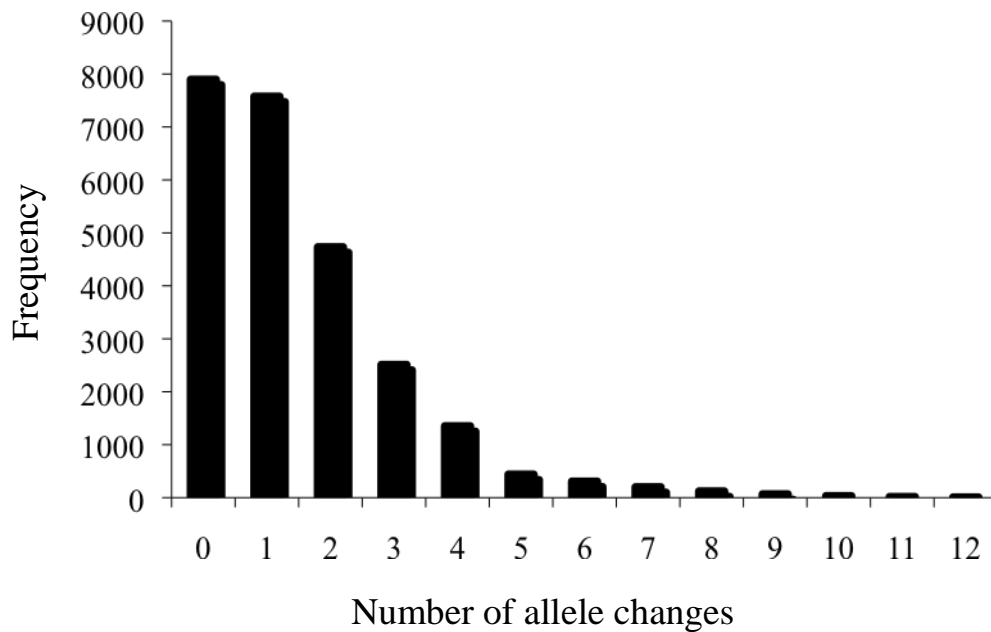


Figure 4B. The mismatch distribution of COI sequences obtained from larvae populations. The bar graph illustrated the uni-modal wave distribution of allele changes in COI sequences. The mean tau value of allele changes was 2.565.

In the adults, the lack of significance of the Harpending's raggedness index also supported the sudden expansion effect. Notably, there were more than more instances of the same haplotype (number of differences among sequences = 0) in larval populations (discussed below) which may have contributed to the raggedness of the larval mismatch. It is also possible that the lower number of sites used in larval comparisons influenced the mismatch shape.

Adult and larval expansion estimates indicated the time since expansion having an older interval in the larval population sample, using the formula $\tau = 2ut$ and mutation rate (μ) of 2.3% per million years.

Using the adult sequences of 359 bp, the mean τ for 10 populations was 2.253, (95% confidence interval: 1.200 – 3.208); and $\theta_0 < 0.001$; and $\theta_1 = 70038.443$; where θ_0 is the mutation parameter estimate and θ_1 the expansion under an infinite allele model of mutation (Rogers 1995). In this case, t was estimated at 0.242 MY (CI: 0.145 – 0.389 MY). The median estimate of t using the larval sample (222 bp) was 0.502 MY ($\tau = 2.565$, and the 95% confidence interval ranged from 1.068 to 4.889 MY). The larval mismatch distribution variance had a higher upper limit ($t = 0.209 – 0.976$ MY) than the adult population mismatch. Sample larvae represent the general population pool and look to have a larger N_e if tau ($\tau = 2.565$) is higher.

Population structure

Pairwise F_{ST} tests and AMOVA indicated no regional structure, but local spatial difference was seen in the comparison of some adult populations. Comparison of megalopae collections over different times revealed varying relative haplotype frequency, which suggests local variance is due to reproductive success rather than regional dispersal.

Population structure analysis considering adults

While there was no indication of large-scale regional differentiation, adult populations showed small-scale regional differentiation (Table 5A). The overall adult pairwise F_{ST} considering 10 populations was 0.090, $p < 0.001$. 64% of the adult pairwise F_{ST} comparisons (23 out of the 36 in Table 4A) were statistically significant ($p < 0.05$). Three adult populations – Dungeness Bay, Astoria, and Eureka – had significantly high pairwise F_{ST} (all F_{ST} scores were below $p < 0.01$) as compared to all other populations.

Table 5A. A pairwise F_{ST} matrix for adult *C. magister* populations. Label abbreviation and dates correspond to Table 3.A. Bold indicates p-values of ≤ 0.05 ; * indicates $p \leq 0.01$; ** indicates $p \leq 0.001$.

DB2009	–									
GH2009	0.108	–								
IW2009	0.177**	0.043	–							
AS2004	0.198**	0.122**	0.068*	–						
EK2005	0.119*	0.085*	0.158**	0.171**	–					
SF2005	0.175**	0.077*	0.021	0.069**	0.121**	–				
HB2005	0.205**	0.090	-0.010	0.081**	0.174**	0.029	–			
ML2005	0.149**	0.052	-0.003	0.088**	0.130**	0.032	0.003	–		
SL2005	0.211**	0.056	-0.003	0.088**	0.157**	0.021	0.017	0.022	–	
FD2009	0.194**	0.129**	0.018	0.068*	0.175**	0.013	-0.008	-0.015	0.040	–

The Mantel and the Exact tests supported AMOVA inference, revealing no structure throughout the range in adult samples. The Mantel test correlating the adult pairwise F_{ST} scores and geographic distance among sites was non significant ($r = 0.059$ $p = 0.289$). The Exact test also could not reject the null hypothesis of a non-random distribution of haplotypes among the samples from along *C. magister*'s range ($p = 0.45$).

Population structure analysis sampling megalopae.

The mean pairwise F_{ST} for the total megalopae population was 0.387 ($p < 0.001$) and ranging from 0.037 to 0.510. The lowest pairwise F_{ST} value was between years 2008 and 2009 (Table 5B). The highest difference was between 2001 and 2008 (Table 5B).

Table 5B. A pairwise F_{ST} matrix for megalopae populations from Coos Bay, OR. Label abbreviation and dates correspond to Table 3.B. Bold indicates p-values of ≤ 0.05 ; * indicates $p \leq 0.01$; ** indicates $p \leq 0.001$.

2001	—			
2003	0.443**	—		
2008	0.510**	0.332**	—	
2009	0.421**	0.207**	0.037*	—

Larval samples were grouped by annual collection time for analysis. Data samples compared pairs of years larvae were collected for the pairwise F_{ST} test (Table 5B). In addition to comparing by years, larvae were tested as separate pairs of monthly groups across each year, and also by weekly groups across years. Although data sets pooled by week and then again by month revealed significant differences, the year-to-year comparisons revealed the greatest differentiation and shown in Table 5B.

AMOVA test for regional differentiation

Different year groups of both adult and megalopae samples had higher Φ_{ST} difference more than even pairs of adult populations sampled at different geographic ends of the 3000 km coast line. The first AMOVA arrangement (Table 6) was conducted to test crab populations from 3 geographic regions indicated in comparison 1 and included adults and larvae, showing no evidence of regional differentiation ($\Phi_{CT} = -0.004$, $p = 0.477$). In comparison 2, geographic structure was not apparent across the 3 large-scale regions when adults only were considered ($\Phi_{CT} = 0.013$, $p = 0.096$) (comparison 2; Table 6). Within these regions, *within-group* (Φ_{SC}) and *inter-population* variance (Φ_{ST}) components were significant (see comparison 2 and 3; Table 6).

Table 6. AMOVA tests of alternative spatial and temporally based comparisons. Comparisons are listed by number for reference. P-values are given in parentheses and bold numbers emphasis significant values.

Comparison	AMOVA test arrangement	Φ_{CT} (p)	Φ_{SC} (p)	Φ_{ST} (p)
1	3 <i>spatial</i> groupings of adults (North FD, DB, GH, IW; Central AS, EK; and South SF, HB, ML, SL) and megalopae from Coos Bay.	-0.004 (0.477)	0.282 (< 0.001)	0.274 (< 0.001)
2	3 <i>spatial</i> groupings of adults only (North, Central, and South).	0.013 (0.096)	0.081 (< 0.001)	0.093 (< 0.001)
3	2 <i>spatial</i> groupings of adults only (North FD, DB, GH, IW, AS; South EK, SF, HB, ML, SL).	-0.008 (0.670)	0.093 (< 0.001)	0.086 (< 0.001)
4	6 <i>temporal</i> groupings of adults and megalopae by net year (2001, 2003, 2004, 2005, 2008, and 2009).	0.212 (0.037)	0.107 (< 0.001)	0.297 (< 0.001)
5	3 <i>temporal</i> groupings of adults only (2004, 2005, and 2009).	0.039 (0.023)	0.065 (< 0.001)	0.102 (< 0.001)
6	4 <i>temporal</i> groupings of megalopae from Coos Bay arranged by month 2001: June, July, August; 2003: June July August; 2008: April, May, June, July, August, September; 2009: April, July August.	0.197 (0.035)	0.618 (< 0.001)	0.694 (< 0.001)

To determine if the spatial structure occurred in a simpler North to South scale rather than 3 regions, adult populations were arranged as two broader regions (comparison 3; Table 6), These also revealed no structuring: $\Phi_{CT} = -0.008$, $p = 0.670$.

AMOVA based comparisons of megalopae sampled at different times

AMOVAs were next conducted across 3 separate arrangements of adults by year, adult and larvae by year, and finally adults and larvae by month (comparison 4, 5, and 6;

Table 6). Differentiation was often observed in comparisons of adults and larvae ($\Phi_{CT} = 0.212$, $p = 0.037$, comparison 4). In comparison 4 (Table 6), megalopae from Coos Bay from the years 2001, 2003, 2008 and 2009 were compared to adults from 2004 (which included only AS), 2005 (which included SF, HB, ML, SL) and 2009 (which included FD, DB, GH, and IW). Temporal variance of megalopae and adults (comparison 4) *within-group* years ($\Phi_{CT} = 0.212$) was similar to comparison 1 (Table 6) spatial *within-group* variance ($\Phi_{SC} = 0.282$).

The highest variance levels in AMOVA occurred in comparing groups of larvae only collected across multiple years (comparison 5 and 6; Table 6). The strongest variance was found in comparison 6 (Table 6) in which megalopae samples were broken down into collection months through out 2001, 2003, 2008, and 2009. An AMOVA across these 4 annual groups showed significant differences among monthly samples ($\Phi_{CT} = 0.197$ $p = 0.035$) as well as *within-group* and *within-population* variance when arranged by months.

The allele frequency difference observed among concentrated groups of larvae could be a result of related larvae rafting together over 6 months as they disperse in the open ocean, completing stage I-V development. This effect may be seen in examining the frequency of private haplotypes (Table 4), as there are a low number of unique haplotypes and mostly common haplotypes are sampled together. Megalopae were light-trap sampled through out an entire month in time periods ranging over the course of about 7 to 14 days, suggesting broods of larvae (characterized by a single private haplotype) are recruiting into estuaries together after months at sea.

Consideration of diversity measures and environmental variables

We examined population diversity measures (h or π) to determine whether they correlate with easily obtained environmental variables. Larvae were not considered for this relationship test, having been sampled in a single location, Coos Bay, OR. Adult genetic diversity (Table 3A) and environmental variables (Table 7, 8, and 9) were tabulated and data was then tested for statistical normality. The Shapiro-Wilk (1965) test indicated that NTU was not normally distributed and was subsequently transformed with natural log for analyses as an environmental variable (Table 9).

Table 7. Reported turbidity measurements (Nephelometric Turbidity Units, averaged) at selected sites with references cited.

Area	Average turbidity (NTU)	Location notes	Geographic density of sampling	Sampling frequency	Collection period and data Source
Fraser Delta, BC, Canada	45	Strait of Georgia, estuary mouth	24 stations	not given; daily average	04/1991-05/1991, (Gregory 1998); 04/1999-01/2001; (Masson 2002)
Dungeness Bay, WA	2	Dungeness River, estuary mouth	4 stations	once a month, for 12 months	11/1999-10/2000; (Sargeant 2000)
Grays Harbor, WA	75	Upper estuaries	5 stations	9 samples over five days	19-21/05/1983; (Kehoe 1985)
Ilwaco, WA	3.5	Bakers Bay	4 randomly positioned samples	4 samples daily	06/1977-03/1978; (Durkin et al. 1979)
Astoria, OR	25	Youngs Bay	4 randomly positioned samples	4 samples a day	06/1977 and 03/1978; (Durkin et al. 1979)
Eureka, CA	8.2	North Humboldt Bay	20 sites along channel	once per month	08/2003-07/2004; (Pinnix et al. 2001)
San Francisco, CA	8.6	San Pablo bay, Central Bay, South Bay	4 sample sites at 3 locations	once a month	10/2000- 01/2001; (Buchanan & Ganju 2001)
Half Moon Bay, CA	3.6	Beachline	One site	2 samples on one day	2009 (date not given); (Lean et al. 2010)
Moss Landing, CA	8	Elkhorn Slough mouth and within 20m of river	two sites	once a month for 12 months	12/1999-09/2001; (Chapin et al. 2004)
Port San Luis, CA	1.5	Along dock at Alvisa bay	1 km from shore at surface and 10 meters depth	every half hour for 4 hours	06/2009-08/2009; (Orric et al. 2009)

Table 8. Physical and genetic measurements from collection sites.

Site	Opening L x W	Depth (m)	Catch in Tons	Velocity (cm ³ /s)	Turbidity (NTU)	Haplotype diversity <i>h</i>	Nucleotide diversity π
Fraser Delta	49.29	62	5599	30	45	0.63	0.003
Dungeness Bay	13.30	1.4	NA	15	2	0.87	0.008
Gray's Harbor	16.85	10	772.5	97.74	75	0.82	0.007
Ilwaco	21.27	4	6.31	87.45	3.5	0.82	0.004
Astoria	11.70	2.5	NA	108.02	25	0.98	0.014
Eureka	0.61	3.6	7216	2	8.2	0.95	0.008
San Francisco	30.79	41	2271	80	8.6	0.84	0.006
Half Moon Bay	27.42	9.1	Included above	25	3.6	0.79	0.004
Moss Landing	21.76	7.6	92.5	25	8	0.86	0.006
Port San Luis	34.24	11	Included above	11	1.5	0.86	0.005

Table 9. Statistical description of Skewedness. Skewedness results of physical and genetic measurements as well as Shapiro-Wilk test of normality with p-value used for correlation. Because the NTU dataset was higher than the skewedness threshold of 2.0, it was logarithmically transformed and retested.

	LxW	Depth	Velocity	lnNTU	<i>h</i>	π	<i>F_s</i>	<i>D</i>
Number of cases	10	10	10	10	10	10	10	10
Variance	184.684	400.806	1,623.700	1.613	0.009	0.001	27.932	0.706
Skewness	0.427	1.920	0.454	0.930	-0.917	1.598	-1.507	1.232
Shapiro-Wilk Statistic	0.982	0.688	0.856	0.892	0.904	0.860	0.875	0.902
Shapiro-Wilk p-value	0.976	0.001	0.069	0.177	0.243	0.077	0.114	0.228

The comparison of environmental variables with genetic diversities revealed several significant linear relationships (Figure 5). Physical dimensions that did not correlate with the genetic data, include lnNTU at $r^2 = -0.219$ for h and $r^2 = 0.183$ for nucleotide diversity. In addition, velocity did not correlate with diversity measures (gene diversity and velocity was $r^2 = 0.136$ and nucleotide diversity and velocity was $r^2 = 0.475$). Opening size (LxW) related with both gene diversity (h) and to a lesser extent nucleotide diversity (π) (simple Pearson's r^2 was -0.840 for h , but $r^2 = -0.484$ for π was not a relationship). When the mean depth of embayments was compared with gene diversity, there was a negative relationship for h (Figure 5; $r^2 = -0.757$) but a lower trend for π ($r^2 = -0.434$).

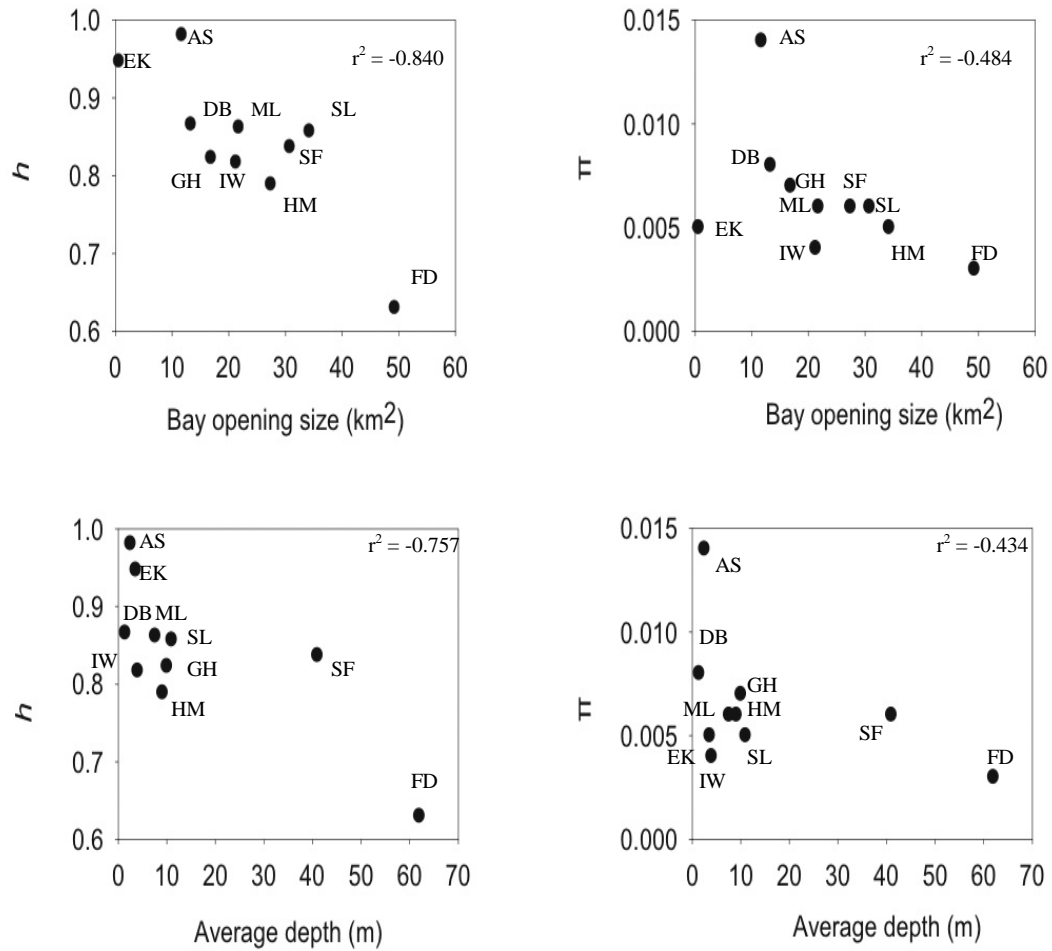


Figure 5. Plots of genetic diversity measures and environmental parameters. Turbidity and velocity data are not shown as there was no obvious relationship. Site abbreviations refer to labels in Table 3A.

A post hoc analysis was conducted to eliminate alternative explanations, using commercial catch numbers, measured by the ton, for each location (graph not shown) with Tajima's D value and gene diversity (h). The commercial catch in tons plotted against h revealed no relationship with $r^2 = 0.038$. While commercial catch in tons plotted against Tajima's D revealed no relationship with $r^2 = 0.0631$. This suggested that the scaling of harvest intensity did not relate to genetic diversity.

Environmental data were also related to adult population pairwise F_{ST} . Only opening dimensions (LxW) showed a significant correlation (Figure 6). In total, 45 pairwise F_{ST} data points were compared, corresponding to pairwise F_{ST} across 10 populations listed in Figure 1. A linear regression procedure indicated a strong negative relationship ($r^2 = -0.67$ $p < 0.001$) between pairwise F_{ST} and LxW (Figure 6). Since the comparisons between pairwise F_{ST} and LxW were not independent measures, the Wilcoxon-signed rank test was used, indicating a significant negative relationship between genetic diversity opening size (Wilcoxon test Z value = 3.622, $p < 0.001$).

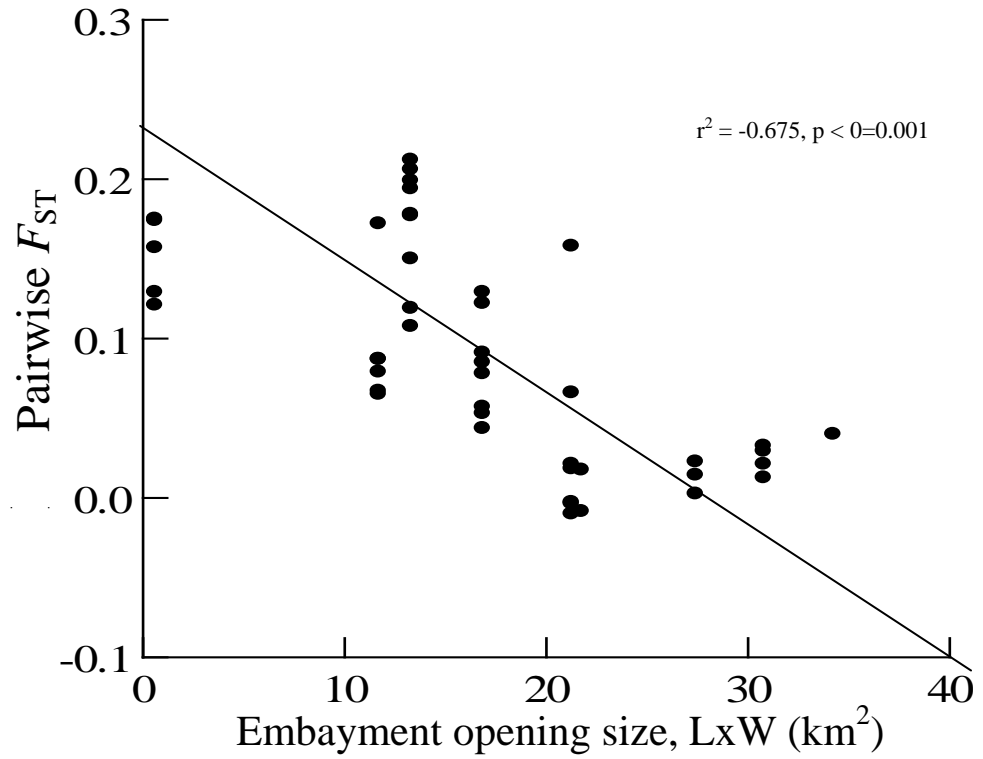


Figure 6. Plot of embayment opening sizes versus F_{ST} . Robust regression was used to calculate a linear relationship for pairwise F_{ST} values against embayment opening size (km^2). Trend line calculated in Systat ($r^2 = -0.675, p < 0.001$). The Shapiro-Wilk Test Z value = 3.622, p-value was < 0.001 .

Canonical correlation comparison of environmental and genetic relationships

Canonical correlation analysis showed strong negative relationships between measures of genetic diversity and environmental measures (Figure 7). The high positive canonical loading for environmental measures (loadings chart in Figure 7) indicated that opening LxW was positively related to canonical scores. In addition, the high positive loadings for gene diversity (h) and for Tajima's D were correlated to opening size canonical scores (Figure 7). The turbidity measure (lnNTU) did relate to genetic variables but was not interpreted because its loading was below a typically accepted threshold (0.424) (data not plotted). The CCA indicated that there was a relationship between the genetic diversity of crabs and the environment in which they obtained. The implication is that coastal terrain has an impact on the genetic measures, in this case, Tajima's D and gene diversity (h).

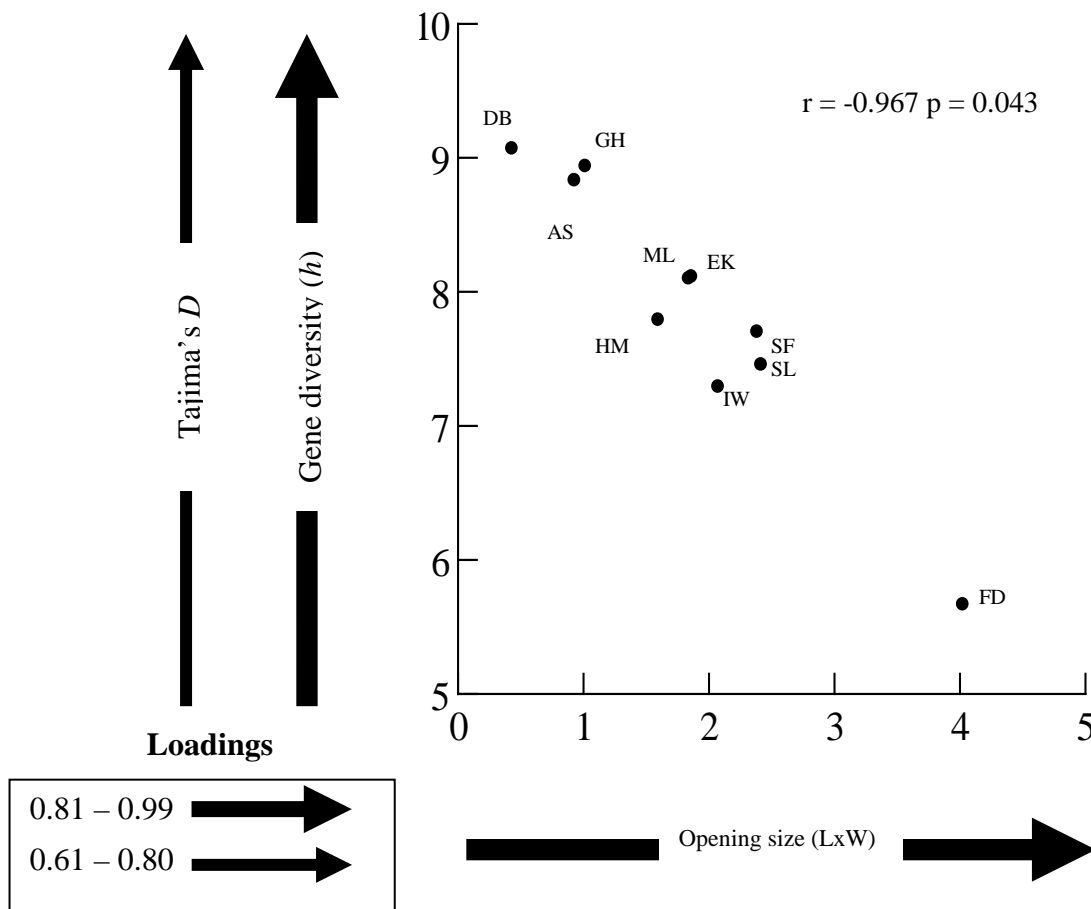


Figure 7. Canonical Correlation Analysis of diversity measures and habitat measures. The plot shows the strongest bivariate correlates. Arrows represent loadings; direction of the arrow indicates the signal of the loading and the thickness of the arrow indicates that absolute value of the loading. Nucleotide diversity π showed the same trend as h , and depth showed the same trend as LxW.

DISCUSSION

Spatial and temporal genetic variation was assessed using mitochondrial DNA from adult and megalopae of Dungeness crabs sampled from Fraser Delta in British Columbia, Washington, Oregon, and California. Megalopae were sampled repeatedly using light traps that were deployed in one location, Coos Bay, OR. Nucleotide (π) and gene diversity (h) for the entire adult dataset were 0.010 and 0.850, COI diversity was uniform and lacking the variance found in adult populations, which is remarkable considering larvae were caught across several years. While this study found no net genetic difference consistent with regional barriers to dispersal, genetic differences were evident in comparing separate years of larval samples (Table 3B) as discussed below.

Given the long distance dispersal of *C. magister* larvae, we expected to find was low to no spatial variation among regions as reported in other studies (Dawson 2001, Petersen 2007, Kelly & Palumbi 2010), and this was consistent with our F_{ST} comparisons of adult populations. Secondly, there was no correlation between geographic distances and genetic distances. Habitat variables, however, were found in some cases to correlate measures of COI diversity. The adult populations sampled in smaller bay tended to have higher diversity (h).

We found a third ecologically significant pattern to diversity. Low diversities of broods of larvae (while warranting further statistical examination) were apparent. This

observation suggests larvae from the same brood aggregate during dispersal remaining grouped even at settlement (following about 2-month pelagic length dispersal).

No apparent broad scale differentiation

In sampling extensively over the range of a species, it may be possible to detect distance-limited dispersal, where genetic distances between demes are correlated with their spatial separation, in other words, an Isolation-by-Distance pattern (Wright 1931, Hellberg 1996, Whitlock & McCauley 1999, Holsinger & Weir 2009, Weersing & Toonen 2009). There was no significant fit to the Isolation-by-Distance model, considering adult populations sampled from San Luis Obispo, CA to Frasier Delta, BC with a coastal length covering 1700 km (Mantel $r = 0.059$, $p = 0.289$) or evidence of discrete genetic breaks.

This study suggests accordingly that high gene flow among adult populations leading homogeneity over a large geographic scale (British Columbia to California). The non-significant Mantel and Exact tests showed no relationship between geographic distance and genetic differentiation among population. In addition, AMOVA indicated that variance among regions was non-significant, suggesting no population structure across the *C. magister* range. Beacham et al. (2008) and Barney (2012) found similar results using COI and microsatellites.

The lack of geographic structure found in this study contrasts with the geographic variation found in the COI gene of *Hemigrapsus oregonensis* (Petersen 2007), a species with 3-4 week larval dispersal. Petersen did suggest a possible genetic dilution in *C.*

magister that we observed in this study (2007). *C. magister* had no obvious barrier to gene flow from British Columbia to California despite some evidence of small scale variance. Similarly to *H. oregonensis* (Petersen 2007), neutrality tests on both adults and megalopae indicated Dungeness crab went through a recent, post-bottleneck population expansion, likely associated with the most recent glacial relaxation, and accordingly, population samples always consisted of many haplotypes.

High genetic variance among larval collections

Small-scale spatial variation was evident. The typical pairwise F_{ST} seen in comparing samples of larvae from Coos Bays by separate years (range: 0.04 - 0.51, mean = 0.33) was about three times higher than pairwise F_{ST} 0 - 0.21, mean = 0.09), as shown in Table 3.

AMOVA revealed variation when considering time-sampled within-group populations (see Φ_{SC} in Table 6 in all 6 comparisons). These temporal variations could be seen when comparing either adult or megalopae samples. An AMOVA comparing larval samples (pooled across months) among 6 non-consecutive years from Coos Bay, showed a difference being retained over the years: $\Phi_{CT} = 0.212$, $p = 0.037$. This suggests that diversity is potentially maintained by annual variance of successful recruits. It appears that the complex diversity of adult populations and lack of broad regional distinction are the result of movements of pulses of larvae. Recruitment over time, in other words, dampens genetic differentiation. Temporal variations could be seen when comparing sequences of adult or megalopae samples. An AMOVA comparing larval

samples (pooled across months) among 6 non-consecutive years from Coos Bay revealed a difference being retained over the years: $\Phi_{CT} = 0.212$, $p = 0.037$. This suggests that diversity is potentially maintained by annual variance of successful recruits.

Inverse correlation among bay size and genetic diversity

Species that are isolated in small populations typically have higher rates of divergence with lower diversity than comparable species of larger populations (as suggested by Ohta 1992, Gillespie 1999, 2001, Woolfit & Bromham 2005). In *C. magister*, the measured COI sequence diversity was correlated to embayment opening sizes. Analyses the relationship between genetic diversity indices such as π , Tajima's D , and Fu's F_S are still not often considered (e.g. see Wares 2009). We found unexpectedly, small bays tended to have greater genetic diversity than large bays.

The relationship between genetic data and hydrologic features are easily accessible means of identifying potentially high diversity in populations. While COI diversity was negatively related to bay size and depth, the size and depth measures were themselves positively correlated. A canonical correlation analysis in this sense was useful for better understanding trends. The CCA considering bay dimensions and genetic diversity measures (π , h , and Tajima's D , and Fu's F_S) supported a strong correlation between bay opening size and gene diversity, h , with a significant correlation also seen between the opening size and D .

The negative correlation between genetic diversity and bay size may be explained a number of ways. Small embayments may allow high survivorship of unique haplotypes

(i.e., through reduced harvesting), thereby increasing gene diversity measures; or perhaps due to imprecise collection methods of adults, one might see an artifact of sampling ocean shelf habitats with varying diversity patterns.

Perhaps recruitment occurs less often in restricted bays leading to a greater spatial genetic patchiness in these areas. We used the opening size (LxW) to approximate restriction of larval movement in and out of an embayment. This restriction is not a direct measure of the bay water's turnover rate, but still may represent the difficulty of *C. magister* entering and exiting a bay. Further studies comparing artificial coastal drifters and water turnover (e.g. Hickey & Banas 2003) might show whether a current based model could predict the genetic variance effect that we measured. Accordingly, the mechanism of how hydrology affects genetic diversity is uncertain, and may warrant a focused sampling regime. We do not always have precise sample locations for the adult specimens examined because some samples were obtained entirely or in part from fishing boats. An aspect of the variance might be due to a collection artifact, that is, small bays may harbor vessels that travel outside the bay to bring in their catch. Just as easily, changes in hydrology could be a reasonable explanation for annual variance among Dungeness crab distribution (Galindo et al. 2006).

The maternally inherited mitochondrial data presented in this project will reflect the female effective populations. We found a stronger correlation between bay size and genetic diversity than between regional fishing pressure and diversity. This suggests that overall no effect of harvesting on diversity was obvious. But determining whether an effect could occur is difficult – there have been no comparison of populations present

prior to commercial harvesting. We observed an unusual genetic signature in the San Francisco populations, where haplotype 2 was not sampled, despite being common elsewhere. Another challenge, which could be addressed in future research, is the need to gain more localized measures of variation in harvest pressure. There may be harvesting related effect in the dataset.

Diversity pattern in the larval pool

There was a tendency of observing larvae that shared haplotypes in a given light trap collection, which suggests that larvae from the same brood often remain aggregated during dispersal, then recruiting together back into estuaries. The data suggest that often single haplotypes (See Table 4 for the year and weeks where only one or two haplotypes are sampled) appear together during the recruitment period resulting in low diversity.

The implication is that a few parents are producing large numbers of offspring and those offspring are maintain kin bonding during 5 to 6 months of open ocean dispersal.

Diversity of larval samples collected together prior to settlement was much lower (2001 $h = 0.583$ while 2009 $h = 0.821$) than in adult populations (mean $h = 0.851$). The relatively low diversity of collected haplotypes in larval samples demonstrates what has been referred to as the “Hedgecock effect” (Waples 1998) – a low diversity of propagules reflecting reproduction by a few mothers in the population. This thesis showed that families of larvae may remain aggregated during 5 months of development. The larval haplotype variation was as low as one haplotype (the pattern reflecting perhaps one mother only) in many samples. Over time however, there was additive variation in the

reproducing females that resulted in high genetic diversity, with the frequency of long distance dispersal events being sufficiently high, preventing regional patterns of differentiation.

There are multiple plausible explanations for genetic variance among crab populations. Post settlement selection and commercial harvesting may have had significant impacts on a population signal after larval settlement. Contrasting these post selective factors, the current study has shown that variance occurs in megalopae prior to settling.

Hedgecock (1994) predicted that adult diversity would have higher diversity than in the surviving offspring and demonstrated that stochastic environmental factors would determine that a fraction of successful broods return each year. Our results demonstrate that the larval populations sampled have lower N_e than adult populations. This pattern is common in other marine species, including crustaceans, echinoderms, and fish (Hedgecock 1994, Gillespie 2001, Turner et al. 2002). Our study supports Hedgecock's prediction of reproductive variance, where the diversity of adult populations might be estimated as functions of temporal variation in the alleles present in recruiting megalopae. We showed a higher diversity in adult populations and lower diversity in larval cohorts. This suggests that over all population diversity might be maintained by the influx of low-diversity sets of larval recruits. Over time, due to the high mortality of offspring year after year, megalopae contribute to a small portion of *C. magister*'s overall diversity.

Variance among years has been identified in the correlation between hydrology and diversity of local populations. On a local level, the narrow and shallow hydrological

features of embayments impact the diversity of adult crabs. The mechanism of hydrological features has been discussed in Roegner et al. (2003), where they suggested the Columbia River greatly increased larval distribution. Accordingly, further studies of the genetic variation patterns in larval populations are needed.

In summary, the diversity data of the COI sequence revealed surprising ecological patterns across the range of *C. magister*. Genetic differences accumulated over time are a more likely explanation of adult diversity than the alternative explanation of geographic distribution. In addition to accumulative variance, local hydrological features may also contribute the adult's variance even between nearby populations.

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